

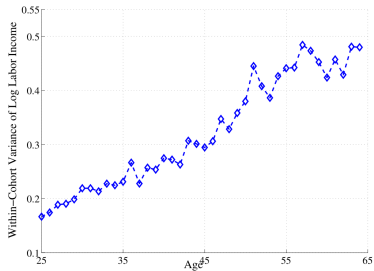
“Inferring Labor Income Risk From Economic Choices: An Indirect Inference Approach”

by Fatih Guvenen and Anthony Smith

October 27, 2010

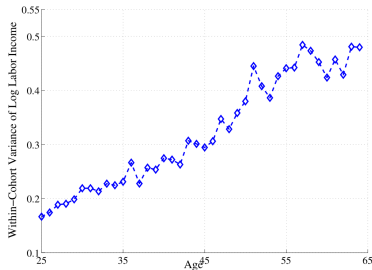
Motivation

- It is well documented that within-cohort dispersion of labor earnings rises substantially over the life cycle.



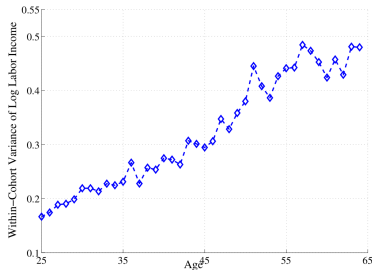
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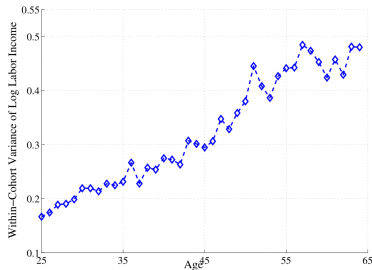
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- There are different interpretations of this pattern, with different policy and welfare implications
- One is that the rise in inequality stems from the accumulation of persistent idiosyncratic shocks.
- The alternative is a systematic fanning out of earnings over time, which can represent uncertainty if individuals have to learn their growth rate.



Overview

- This paper will assume that individuals face the following income process

$$y_t^i = \underbrace{g(t, \text{observables}, \dots)}_{\text{common life cycle component}} + \underbrace{\alpha^i + \beta^i t}_{\text{profile heterogeneity}} + \underbrace{z_t^i + \epsilon_t^i}_{\text{stochastic component}} \quad (1)$$

where $z_t^i = \rho z_{t-1}^i + \eta_t^i$; η_t^i and ϵ_t^i are mean zero innovations.

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- Thus they create a life cycle consumption-savings model with CRRA utility, borrowing constraints and retirement system.
- They will estimate this model using indirect inference on the PSID data set with non-durable consumption data imputed from the CEX.
- The auxiliary model will be motivated by the model above under quadratic utility and no borrowing constraints

The Learning Problem

- The assume that $\beta^i = \beta_k^i + \beta_u^i$ where β_k^i is observed and $\sigma_\beta^2 = \sigma_{\beta_k}^2 + \sigma_{\beta_u}^2$. Let $\lambda \equiv \frac{\sigma_{\beta_u}^2}{\sigma_\beta^2}$.

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- Learning β^i can then be written as a Kalman Filter problem with state equation

$$\begin{bmatrix} \beta^i \\ z_{t+1}^i \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & \rho \end{bmatrix} \begin{bmatrix} \beta^i \\ z_t^i \end{bmatrix} + \begin{bmatrix} 0 \\ \eta_{t+1}^i \end{bmatrix}$$

With a second observation equation

$$\tilde{y}_t^i \equiv y_t^i - \alpha^i - \epsilon_t^i = \begin{bmatrix} t & 1 \end{bmatrix} \begin{bmatrix} \beta^i \\ z_t^i \end{bmatrix}$$

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- For future reference we will define $\hat{\xi}_t^i = \tilde{y}_t^i - \mathbb{E}_{t-1}(\tilde{y}_t^i)$

The Linear-Quadratic Framework

- The consumption-savings problem in the quadratic utility framework can be written as

$$V_t^i(a_t^i, \hat{\beta}_t^i, \hat{z}_t^i) = \max_{C_t^i, a_{t+1}^i} \left\{ -(C_t - C^*)^2 + \frac{1}{1+r} \mathbb{E}_t \left[V_{t+1}^i(a_{t+1}^i, \hat{\beta}_{t+1}^i, \hat{z}_{t+1}^i) \right] \right\}$$

subject to

$$C_t^i + a_{t+1}^i = (1+r)a_t^i + Y_t^i$$

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- One can see that the optimal consumption satisfies

$$\Delta C_t = \varphi_t \left[\sum_{s=0}^{T-t} \left(\frac{1}{1+r} \right)^s (\mathbb{E}_t - \mathbb{E}_{t-1}) Y_{t+s} \right]$$

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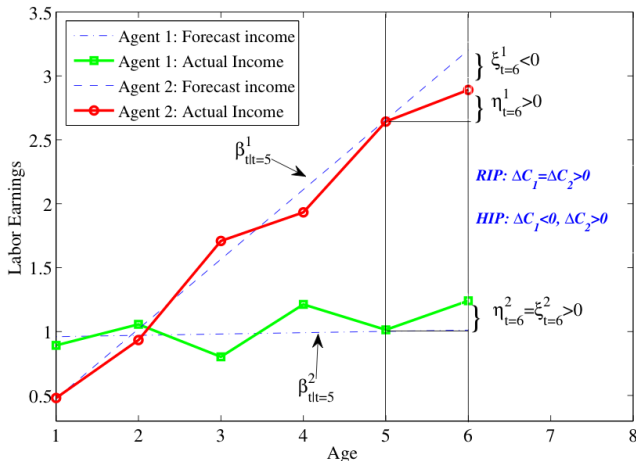
$$\Delta C_t = \varphi_t \left[\sum_{s=0}^{T-t} \left(\frac{1}{1+r} \right)^s (\mathbb{E}_t - \mathbb{E}_{t-1}) Y_{t+s} \right]$$

- All new information comes from $\hat{\xi}_t^i$, thus one can obtain $\Delta C_t = \Pi_t \times \hat{\xi}_t^i$ and derive the level of optimal consumption

$$C_t^i = \varphi_t w_t^i + \frac{1}{1+r} \Phi_{t+1} \hat{\beta}_t^i + \frac{1}{1+r} \rho \Psi_{t+1} \hat{z}_t^i$$

Some Examples

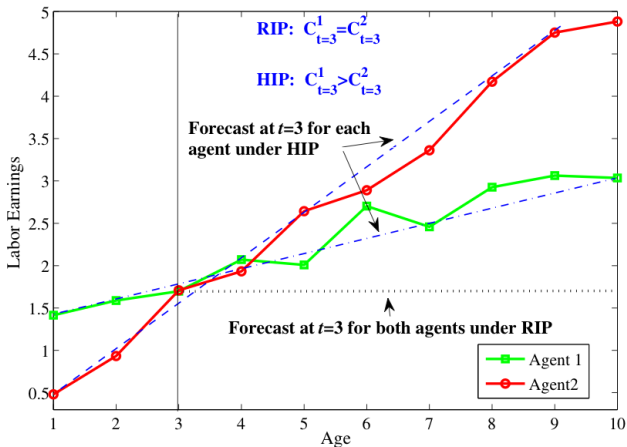
From this framework one can see that individuals with different beliefs about their growth rate β will respond differently to the save earnings innovation. This will help determine beliefs in the estimation



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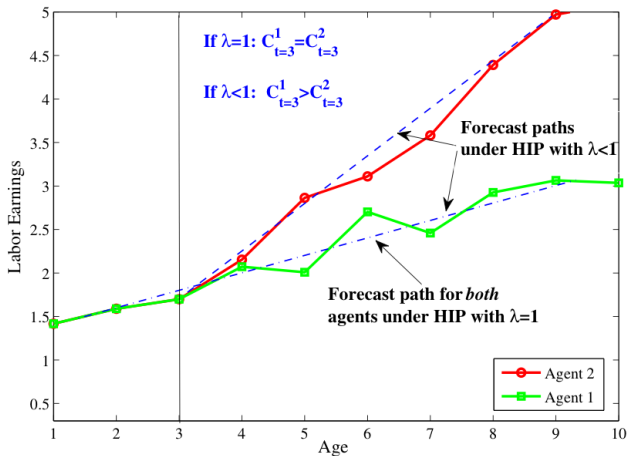
Proposition

Controlling for and individuals income and assets, the consumption level is an increasing function of the individuals beliefs $\hat{\beta}_t^i$.



Some Examples

Consumption decisions can also be informative about the degree of private information, λ .



Full Consumption Savings Model

- Agents live T periods and retire at date $R < T$. The Dynamic Program for the full consumption savings model is

$$V_t^i(w_t^i, \hat{\beta}_t^i, \hat{z}_t^i; \alpha^i) = \max_{C_t^i, a_{t+1}^i} \left\{ \frac{(C_t^i)^{1-\phi}}{1-\phi} + \delta \mathbb{E}_t \left[V_{t+1}^i(w_{t+1}^i, \hat{\beta}_{t+1}^i, \hat{z}_{t+1}^i; \alpha^i) \right] \right\}$$

subject to

$$C_t^i + a_{t+1}^i = w_t^i$$

$$w_t^i = (1+r)a_t^i + Y_t^i$$

$$Y_t^i = \underline{Y} + \exp(\bar{y}_t + \alpha^i + \beta^i t + z_t^i + \epsilon_t^i)$$

$$a_{t+1}^i \geq \underline{a}_t, \text{ and Kalman Recursions}$$

\bar{y}_t is the population average of log income at age t .

- For retirement we have, where $Y^i = \Phi(Y_R^i, \bar{Y})$

$$V_t^i(w_t^i; Y_R^i) = \max_{C_t^i, a_{t+1}^i} \left[\frac{(C_t^i)^{1-\phi}}{1-\phi} + \delta V_{t+1}^i(w_{t+1}^i; Y_R^i) \right]$$

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- They use Blundell, Pistaferri and Preston method for imputing consumption in the PSID using food demand estimates from the CEX.
- The basic approach involves estimating a demand system for food consumption as a function of non-durable expenditures, demographic variables, and relative prices
- They also extend the method to cover the period from 1968 to 1992 (instead of 1980 - 1992).
- They convert the Data to Per-Adult Equivalent Units by regressing each variable on family size, a race dummy a region dummy, a dummy indicating whether the head is employed, a dummy indicating residence in a large city, and a set of cohort dummies. Use the residuals in the analysis below.

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- Thus they use the following equations for the auxiliary model

$$c_t = a_0 + a_1 y_{t-1} + a_2 y_{t-2} + a_3 y_{t+1} + a_4 y_{t+2} + a_5 \bar{y}_{1,t-3} + a_6 \bar{y}_{t+3,R} \\ + a_7 \Delta y_{1,t-3} + a_8 \Delta y_{t+3,R} + a_9 c_{t-1} + a_{10} c_{t-2} + a_{11} c_{t+1} + a_{12} c_{t+2} + \epsilon_t^c$$

and

$$y_t = b_0 + b_1 y_{t-1} + a_2 y_{t-2} + b_3 y_{t+1} + b_4 y_{t+2} + b_5 \bar{y}_{1,t-3} + b_6 \bar{y}_{t+3,R} \\ + b_7 \Delta y_{1,t-3} + b_8 \Delta y_{t+3,R} + \epsilon_t^y$$

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- They also divide the population into 2 groups: between 25 and 38, and between 39 and 55. Allow the coefficients to vary across these two groups.

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- They compute the Wealth to income ratio in 1984 and 1989, average and add the additional moment

$$10 \times (WY_{PSID} - WY_{Model})^2$$

to the objective function.

Estimation

- Used the following objective function for the data

$$\mathcal{L} = |\Sigma|^{-J/2} \exp \left(-\frac{1}{2} \sum_i \sum_t \epsilon_t^{i,Data} \Sigma^{-1} (\epsilon_t^{i,Data})' \right)$$

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$$\hat{\theta} = \operatorname{argmin}_{\theta} \left[\mathcal{L}(\hat{\mathbf{a}}, \hat{\mathbf{b}}, \hat{\Sigma}, PSID) - \mathcal{L}(\tilde{\mathbf{a}}(\theta), \tilde{\mathbf{b}}(\theta), \tilde{\Sigma}(\theta), PSID) \right]$$

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- Use Monte-Carlo to study small sample bias and identification. Find strong evidence for local identification.

Estimation results

- Note that lambda is estimated to be around 0.345 which means that $\frac{\sigma_{\beta,k}^2}{\sigma_{\beta}^2} = 1 - \lambda^2 = .880$.

Baseline	
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σ_{α}	0.284 (0.027)
σ_{β}	1.852 (0.188)
$corr_{\alpha\beta}$	-0.162 (0.151)
ρ	0.754 (0.025)
σ_{η}	0.196 (0.005)
σ_{ε}	0.004 (0.021)
λ	0.345 (0.074)
δ	0.950 (0.001)
ψ	0.874 (0.083)
σ_y	0.147 (0.007)
σ_c	0.356 (0.002)
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- They also did some analysis on the goodness of fit of the model. Over 100 repetitions of the objective function the test statistic has a mean of 2.29 and a maximum value of 4.03. The PSID data yielded a value of 26.37.

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- Note that lambda is estimated to be around 0.345 which means that $\frac{\sigma_{\beta,k}^2}{\sigma_{\beta}^2} = 1 - \lambda^2 = .880$.
- Also estimated discount factor to some precision. This, however, is condition on the fixed value of ϕ . δ is quite sensitive to the preset value of risk aversion.
- They also did some analysis on the goodness of fit of the model. Over 100 repetitions of the objective function the test statistic has a mean of 2.29 and a maximum value of 4.03. The PSID data yielded a value of 26.37.
- This is because the model falls short in matching the auxiliary consumption model.

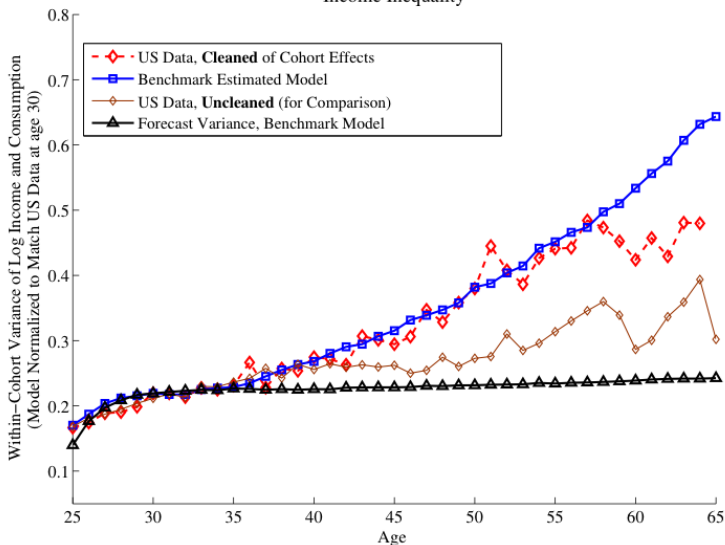
Baseline	
	(1)
σ_{α}	0.284 (0.027)
σ_{β}	1.852 (0.188)
$corr_{\alpha\beta}$	-0.162 (0.151)
ρ	0.754 (0.025)
σ_{η}	0.196 (0.005)
σ_{ε}	0.004 (0.021)
λ	0.345 (0.074)
δ	0.950 (0.001)
ψ	0.874 (0.083)
σ_y	0.147 (0.007)
σ_c	0.356 (0.002)
σ_{c_0}	0.428 (0.009)
max % constrained..	12.7%
.. at age	30
\underline{a}_0 / mean income	0.33
\underline{a}_{55} / mean income	0.53

Goodness of Fit

	<i>constant</i>	y_{t-1}	y_{t-2}	y_{t+1}	y_{t+2}	$\bar{y}_{1,t-3}$	$\bar{y}_{t+3,T}$	$\Delta y_{1,t-3}$	$\Delta y_{t+3,T}$	c_{t-1}	c_{t-2}	c_{t+1}	c_{t+2}
PANEL A: INCOME EQUATION													
<i>Young Group:</i>													
(1) Data	-0.036 ^a	0.346	0.360	0.077	0.097	-0.098	0.150	0.037	-0.022				
(2) Model	-0.026 ^{††b}	0.337	0.357	0.093	0.102	-0.081	0.124	-0.016 ^{††}	-0.021				
<i>Middle-Age Group:</i>													
(3) Data	0.006 ^{**}	0.418	0.358	0.111	0.093	-0.027	0.043 ^{**}	0.031	0.028				
(4) Model	-0.001 ^{††}	0.401	0.375	0.122	0.105	-0.038	0.022	0.012	0.045				
PANEL B: CONSUMPTION EQUATION													
<i>Young Group:</i>													
(5) Data	-0.007	0.108	0.042 [*]	-0.023	-0.005	-0.045 [*]	-0.017	0.030	-0.002	0.248	0.262	0.178	0.175
(6) Model	-0.021 ^{††}	0.105	0.079 [†]	-0.022	-0.039	-0.079	0.025	0.001	0.013	0.208 ^{††}	0.205 ^{†††}	0.248 ^{†††}	0.229 ^{†††}
<i>Middle-Age Group:</i>													
(7) Data	-0.004	0.136	0.046 ^{**}	-0.014	-0.040 [*]	-0.082 ^{**}	0.012	0.030	0.028	0.270	0.260	0.177	0.187
(8) Model	0.007 ^{††}	0.103	0.041	0.009	-0.044	0.050 ^{†††}	-0.042 ^{††}	0.027	0.041	0.197 ^{†††}	0.211 ^{††}	0.224 ^{††}	0.257 ^{†††}

Implications

Income Inequality



Risk Aversion and Time Preference

- Cannot separately identify δ and ϕ , using Monte Carlo and 200 observations find a correlation of the parameters to be -0.88. Suggests only a combination is identified.
- They do some sensitivity analysis and find that one can either fix δ or ϕ , but fixing both will create severe biases. Finally they find that the main role of the Wealth to Income moment is to pin down δ .