

Normalization in Econometrics

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What do we mean by Normalization?

Restricting the parameter space to achieve identification

Example:

- ▶ toy model:

$$y_t = \sigma \varepsilon_t \quad (1)$$

with $\varepsilon_t \sim N(0, 1)$

Log-likelihood is given by:

$$\log f(y; \sigma) = -T/2 \log(2\pi) - T/2 \log(\sigma^2) - \sum_{t=1}^T y_t^2 / (2\sigma^2) \quad (2)$$

Here the 'sensible' normalization would be $\sigma > 0$

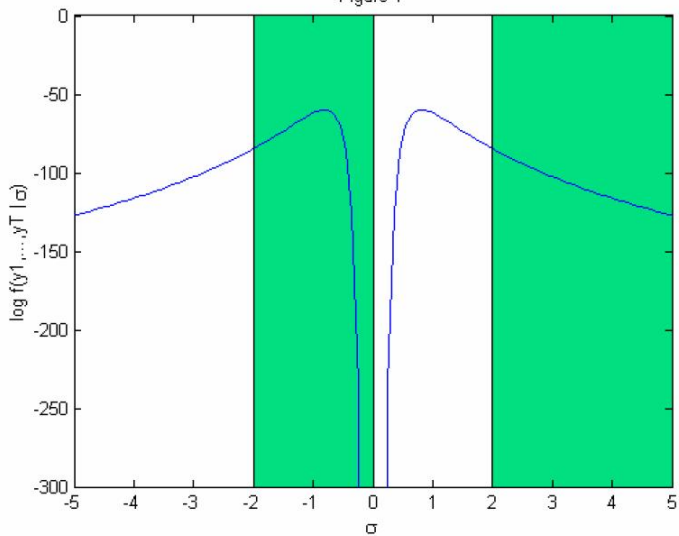
- ▶ Note: this paper only deals with normalization issues in likelihood-based inference

What is the Problem?

Picking a certain normalization not only amounts to 'picking' a point estimate, but also what area of the parameter space is used to compute confidence intervals.

A poor normalization can lead to multimodal distributions, disjoint confidence intervals, and very misleading characterizations of the true statistical uncertainty

Figure 1



Some Definitions

Let $\theta \in \mathcal{R}^k$ be the parameter vector of interest and $f(y; \theta)$ the likelihood function

- ▶ two parameter vectors θ_1 and θ_2 are observationally equivalent if

$$f(y; \theta_1) = f(y; \theta_2) \forall y \quad (3)$$

- ▶ the model is said to be locally identified at θ_0 if there exists an open neighborhood around θ_0 containing no other value of θ that is observationally equivalent to θ_0

The Identification Principle

There will usually be loci along which the model is locally unidentified because the information matrix is singular or the likelihood is 0 ($\sigma = 0$ in the previous example). The following *identification principle* rules out 'bad' normalizations:

The set of admissible parameter values should be chosen in such a way that the loci along which the model is locally unidentified or the likelihood is 0 form the boundary of this set.

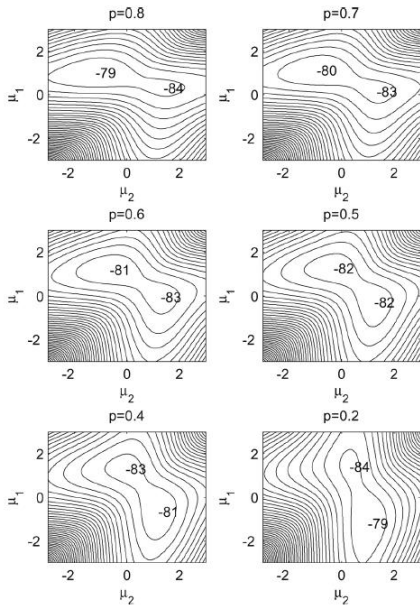
Example 1: Mixture of Normals

$$f(y_t; \mu_1, \mu_2, p) = \frac{p}{\sqrt{2\pi}} \exp\left(\frac{-(y_t - \mu_1)^2}{2}\right) + \frac{1-p}{\sqrt{2\pi}} \exp\left(\frac{-(y_t - \mu_2)^2}{2}\right) \quad (4)$$

Two possible normalization schemes have appeared in the literature:

- ▶ $p > 0.5$
- ▶ $\mu_1 > \mu_2$

Example: 50 observations, $\mu_1 = 1, \mu_2 = -1$ and $p = 0.8$



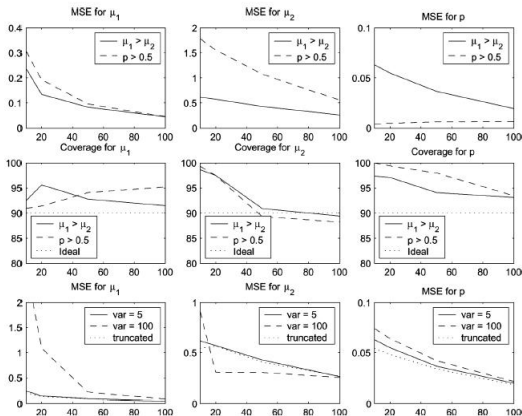


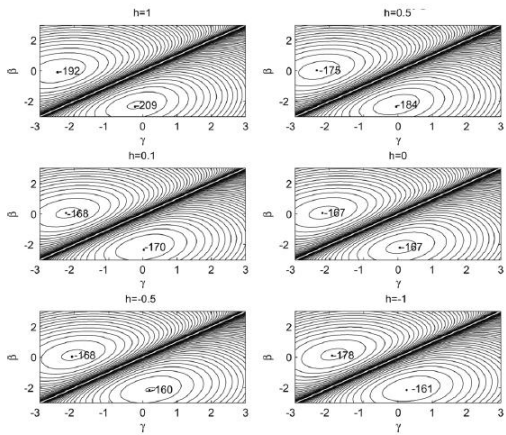
FIGURE 2 Performance of estimators with sample size on the horizontal axis. Top row: average squared difference between posterior mean and true value for indicated parameter when normalized by $\mu_1 > \mu_2$ (solid line) or by $p > .5$ (dashed line). First column: performance for estimate of μ_1 ; second column: performance for estimate of μ_2 ; third column: performance for estimate of p . Second row: ninety-percent coverage probabilities for indicated parameter when normalized by $\mu_1 > \mu_2$ (solid line) or by $p > .5$ (dashed line); dotted line gives nominal 90% goal. Third row: average squared difference between posterior mean and true value for indicated parameter when normalized by $\mu_1 > \mu_2$ for three different priors. Solid line: prior variance = 5; dashed line: prior variance = 100; dotted line: truncated Normal prior.

Example 2: A Structural VAR model

$$B_0 y_t = k + \sum_{j=1}^p B_j y_{t-j} + u_t \quad (5)$$

with $u_t \sim N(0, D^2)$, D diagonal

$$B_0 = \begin{pmatrix} 1 & -\beta & 0 \\ 1 & -\gamma & -h \\ 0 & 0 & 1 \end{pmatrix} \quad (6)$$



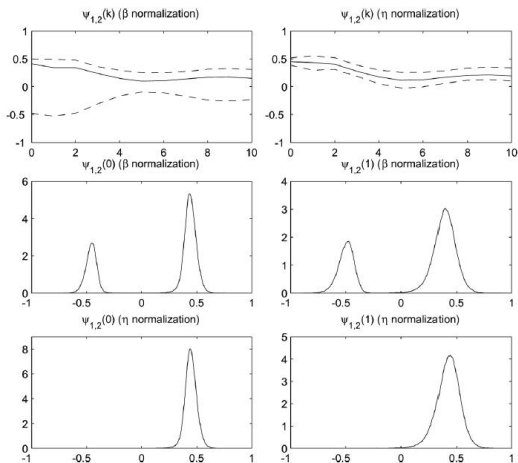


FIGURE 4 Top row: impulse-response function and 90% confidence interval for the effect of a one standard deviation increase in quantity demanded on the price k periods later under the β -normalization (left panel) and η -normalization (second panel). Second row: posterior density for the $k = 0$ (left panel) and $k = 1$ (right panel) values of the impulse-response function plotted in the upper left panel. Third row: posterior density for the $k = 0$ and $k = 1$ values of the impulse-response function plotted in the upper right panel.

Conclusion

- ▶ careless normalization can lead to misleading statistical inference
- ▶ check different normalizations
- ▶ algorithm for structural VARs is available (Waggoner & Zha 2003)